MOMENT-GENERATING FUNCTIONS

1. Demonstrate how the moments of a random variable x may be obtained from its moment generating function by showing that the rth derivative of $E(e^{xt})$ with respect to t gives the value of $E(x^r)$ at the point where t=0.

Show that the moment generating function of the Poisson p.d.f. $f(x) = e^{-\mu}\mu^x/x!$; $x \in \{0, 1, 2, ...\}$ is given by $M(x, t) = \exp\{-\mu\}\exp\{\mu e^t\}$, and thence find the mean and the variance.

2. Demonstrate how the moments of a random variable may be obtained from the derivatives in respect of t of the function $M(t) = E\{\exp(xt)\}$.

If x = 1, 2, 3, ... has the geometric distribution $f(x) = pq^{x-1}$, where q = 1 - p, show that the moment generating function is

$$M(t) = \frac{pe^t}{1 - qe^t}.$$

Find E(x).

Answer: The moment generating function of x is

$$M(t) = \sum_{x=1}^{\infty} e^{xt} p q^{x-1} = \frac{p}{q} \sum_{x=1}^{\infty} (qe^t)^x$$
$$= pe^t \sum_{x=0}^{\infty} (qe^t)^x = \frac{pe^t}{1 - qe^t}.$$

To find E(x), we may use the quotient rule to differentiate the expression M(t) with respect to t. This gives

$$\frac{dM(t)}{dt} = \frac{(1 - qe^t)pe^t - pe^t(-qe^t)}{(1 - qe^t)^2}.$$

Setting t = 0 gives E(x) = 1/p.

3. Let x_i ; i = 1, ..., n be a set of independent and identically distributed random variables. If the moment generating function of x_i is $M(x_i, t) = E\{\exp(x_i t)\}$ for all i, find the moment generating function for $y = \sum x_i$.

Find the moment generating function of a random variable $x_i = 0, 1$ whose probability density function if $f(x_i) = (1 - p)^{1 - x_i} p^{x_i}$, and thence find the moment generating function of $y = \sum x_i$. Find E(y) and V(y).

EXERCISES IN STATISTICS

4. Demonstrate how the moments of a random variable x may be obtained from the derivatives in respect of t of the function $M(x,t) = E(\exp\{xt\})$ If $x \in \{1,2,3...\}$ has the geometric distribution $f(x) = pq^{x-1}$ where q = 1 - p, show that the moment generating function is

$$M(x,t) = \frac{pe^t}{1 - qe^t},$$

and thence find E(x).

5. Demonstrate how the moments of a random variable x—if they exist—may be obtained from its moment generating function by showing that the rth derivative of $E(e^{xt})$ with respect to t, evaluated at the point t=0, gives the value of $E(x^r)$.

Find the moment generating function of $x \sim f(x) = 1$, where 0 < x < 1, and thereby confirm that $E(x) = \frac{1}{2}$ and $V(x) = \frac{1}{12}$.

6.. Demonstrate how the moments of a random variable x may be obtained from its moment generating function by showing that the rth derivative of $E(e^{xt})$ with respect to t gives the value of $E(x^r)$ at the point where t=0.

Demonstrate that the moment generating function of a sum of independent variables is the product of their individual moment generating functions.

Find the moment generating function of the point binomial

$$f(x;p) = p^x (1-p)^{1-x}$$

where x = 0, 1. What is the relationship between this and the m.g.f. of the binomial distribution?

Find the variance of $x_1 + x_2$ when $x_1 \sim f(p_1 = 0.25)$ and $x_2 \sim f(p_2 = 0.75)$ are independent point binomials. (MATHSTATS 97)

Answer. It is straightforward to show that $V(x_1) = V(x_2) = pq$ where $p = p_1$ and $q = p_2$ are used for ease of notation. Since x_1, x_2 are statistically independent, it follows that $V(x_1 + x_2) = V(x_1) + V(x_2) = 2pq$.

Alternatively, we may consider the moment generating functions of the two variables x_1 and x_2 which are respectively

$$M_1 = (p + qe^t)$$
 and $M_2 = (q + pe^t)$.

The moment generating function of their sum $y = x_1 + x_2$ is

$$M = M_1 M_2 = (p + qe^t)(q + pe^t) = pq + q^2 e^t + p^2 e^t + pqe^{2t}.$$

MOMENT-GENERATING FUNCTIONS

Its first and second derivatives are

$$\frac{dM}{dt} = q^2 e^t + p^2 e^t + 2pqe^{2t},$$

$$\frac{d^2M}{dt^2} = q^2e^t + p^2e^t + 4pqe^{2t};$$

and setting t = 0 gives the following moments:

$$E(y) = q^{2} + p^{2} + 2pq = 1,$$

$$E(y^{2}) = q^{2} + p^{2} + 4pq = 1 + 2pq.$$

It follows that

$$V(y) = E(y^2) - \{E(y^2)\} = 2pq = 0.375$$

7. Find the moment generating function of $x \sim f(x) = 1$, where 0 < x < 1, and thereby confirm that $E(x) = \frac{1}{2}$ and $V(x) = \frac{1}{12}$.

Answer: The moment generating function is

$$M(x,t) = E(e^{xt}) = \int_0^1 e^{xt} dx$$
$$= \left[\frac{e^{xt}}{t}\right]_0^1 = \frac{e^t}{t} - \frac{1}{t}.$$

But

$$e^{t} = \frac{t^{0}}{0!} + \frac{t}{1!} + \frac{t^{2}}{2!} + \frac{t^{3}}{3!} + \cdots,$$

SO

$$M(x,t) = \left[\frac{1}{t} + 1 + \frac{t}{2!} + \frac{t^2}{3!} + \frac{t^3}{4!} + \cdots\right] - \frac{1}{t}$$
$$= 1 + \frac{t}{2} + \frac{t^2}{6} + \frac{t^3}{24} + \cdots$$

By the process of differentiating M(x,t) with respect to t and the setting t=0, we get

$$E(x) = \frac{\partial M(x,t)}{\partial t} \Big|_{t=0} = \left[\frac{1}{2} + \frac{2t}{3!} + \frac{3t^2}{4!} + \cdots \right]_{t=0} = \frac{1}{2},$$

$$E(x^2) = \frac{\partial^2 M(x,t)}{\partial t^2} \bigg|_{t=0} = \left[\frac{2}{3!} + \frac{6t}{4!} + \cdots \right]_{t=0} = \frac{1}{3}.$$

EXERCISES IN STATISTICS

Combining these results gives

$$V(x) = E(x^2) - \{E(x)\}^2 = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}.$$

8. Find the moment generating function of $x \sim f(x) = ae^{-ax}; x \geq 0$.

Answer: The moment generating function is

$$M(x,t) = E(e^{xt}) = \int_0^\infty e^{xt} a e^{-ax} dx = \int_0^\infty a e^{-x(a-t)} dx$$
$$= \left[\frac{-ae^{-x(a-t)}}{a-t} \right]_0^\infty = \left[\frac{a}{a-t} \right] = \frac{1}{1-t/a}.$$

9. Prove that $x \sim f(x) = xe^{-x}$; $x \geq 0$ has a moment generating function of $1/(1-t)^2$. Hint: Use the change of variable technique to integrate with respect to w = x(1-t) instead of x.

Answer: The moment generating function is

$$M(x,t) = E(e^{xt}) = \int_0^\infty e^{xt} x e^{-x} dx = \int_0^\infty x e^{-x(1-t)} dx.$$

Define w = x(1-t). Then

$$x = \frac{w}{1-t}$$
 and $\frac{dx}{dw} = \frac{1}{1-t}$.

The change of variable technique indicates that

$$\int g(x)dx = \int g\{x(w)\}\frac{dx}{dw},$$

where $g(x) = xe^{-x(1-t)}$. Thus we find that

$$M(x,t) = \int_0^\infty \frac{w}{1-t} e^{-w} \frac{1}{1-t} dw$$
$$= \frac{1}{(1-t)^2} \int_0^\infty w e^{-w} dw = \frac{1}{(1-t)^2}.$$

Here the value of the final integral is unity, since the expression we^{-w} , which is to be found under the integral sign, has the same form as the p.d.f. of x.

MOMENT-GENERATING FUNCTIONS

To demonstrate directly that the value is unity, we can use the technique of integrating by parts. The formula is

$$\int u \frac{dv}{dx} dx = uv - \int v \frac{du}{dx} dx.$$

Within the expression we^{-w} , we take w = u and $e^{-w} = dv/dw$. Then we get

$$\int_0^\infty w e^{-w} dw = \left\{ \left[-w e^{-w} \right]_0^\infty + \int_0^\infty e^{-w} dw \right\}$$
$$= \int_0^\infty e^{-w} dw = \left[-e^{-w} \right]_0^\infty = 1.$$

10. Using the theorem that the moment generating function of a sum of independent variables is the product of their individual moment generating functions, find the m.g.f. of $x_1 \sim e^{-x_1}$ and $x_2 \sim e^{-x_2}$ when $x_1, x_2 \geq 0$ are independent. Can you identify the p.d.f. of $f(x_1 + x_2)$ from this m.g.f.?

Answer: If x_1 and x_2 are independent, then their joint p.d.f. can be written as $f(x_1, x_2) = f(x_1)f(x_2)$; and it follows that

$$M(x_1 + x_2, t) = \int_{x_2} \int_{x_1} e^{(x_1 + x_2)t} f(x_1, x_2) dx_1 dx_2$$

=
$$\int_{x_1} e^{x_1} f(x_1) dx_1 \int_{x_2} e^{x_2} f(x_2) dx_2 = M(x_1, t) M(x_2, t),$$

or simply that $M(x_1 + x_2, t) = M(x_1, t)M(x_2, t)$. If x_1, x_2 are independent with $x_1 \sim e^{-x_1}$ and $x_2 \sim e^{-x_2}$, then

$$M(x_1,t) = M(x_2,t) = \frac{1}{1-t}$$
 and $M(x_1 + x_2,t) = \frac{1}{(1-t)^2}$.

But, according to the answer to question (3), this implies that

$$f(x_1 + x_2) = (x_1 + x_2)e^{-(x_1 + x_2)}.$$

11. Find the moment generating function of the point binomial

$$f(x) = p^x (1 - p)^{1 - x}$$

where x = 0, 1. What is the relationship between this and the m.g.f. of the binomial distribution?

EXERCISES IN STATISTICS

Answer: If $f(x) = p^{x}(1-p)^{1-x}$ with x = 0, 1, then

$$M(x,t) = \sum_{x=0,1} e^{xt} f(x) = e^{0} p^{0} (1-p) + e^{t} p (1-p)^{0}$$
$$= (1-p) + pe^{t} = q + pe^{t}.$$

But the binomial outcome $z = \sum_{i=1}^{n} x_i$ is the sum of n inependent point-binomial outcomes; so it follows that the binomial m.g.f. is

$$M(z,t) = (q + pe^t)^n.$$